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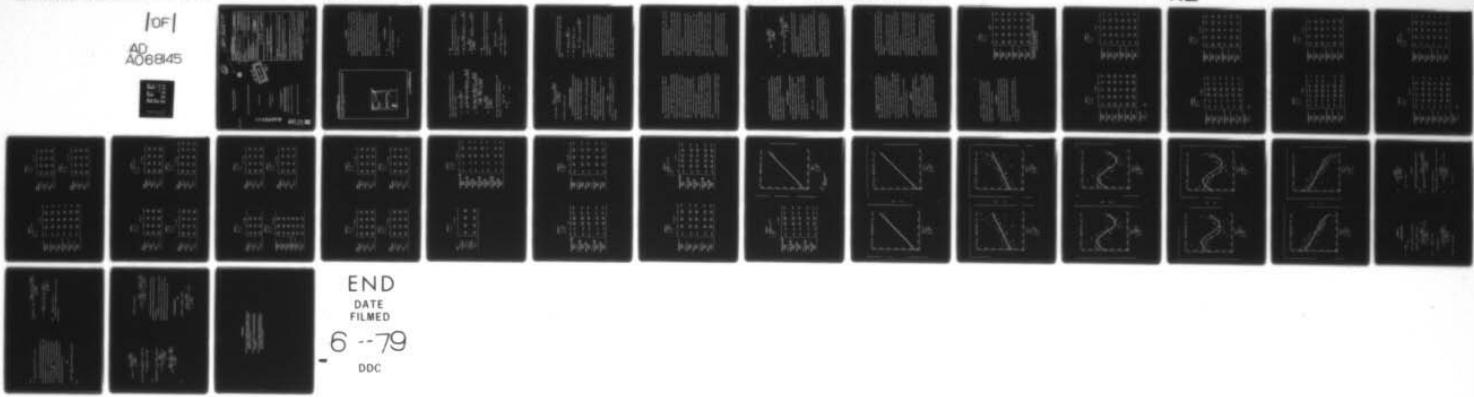
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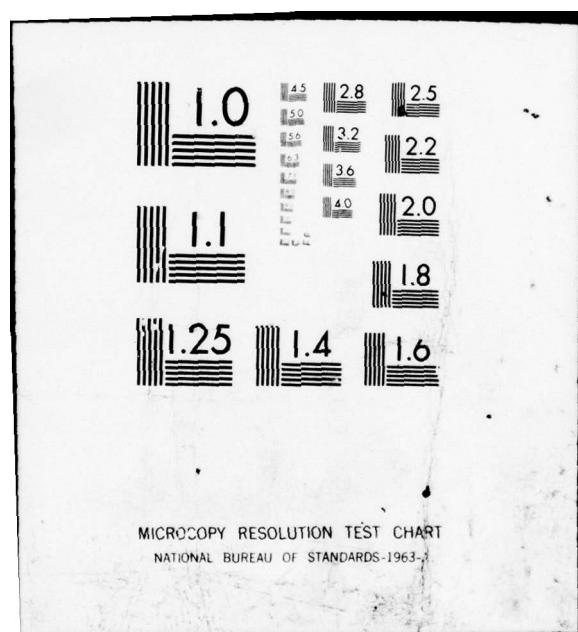
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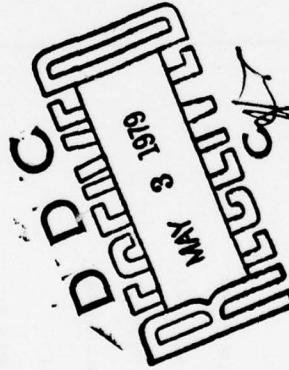
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EMPIRICAL STUDY OF QUANTILE REGRESSION ESTIMATORS



By Michael Meizer

Technical Report No. A-2

February 1979

Texas A & M Research Foundation  
Project No. 3861

"Maximum Robust Likelihood Estimation and  
Non-parametric Statistical Data Modeling"  
Sponsored by the U.S. Army Research Office

Professor Emanuel Parzen, Principal Investigator

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Given bivariate data  $((X_i, Y_i), i = 1, \dots, n)$  one would like a technique

for estimating the regression function  $r(x) = E(Y|X)$ , which is applicable for

a wide range of functions,  $r(x)$ . Further, one would like a technique with

good precision over the full range of values for  $X$ .

Techniques of quantile regression based on the approach developed by

Parzen (1977), and discussed by Carni Meier (1978), will be compared to each.

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other and to Benedetti (1975) in the case where  $X$  is restricted to the unit interval. Further, the quantile regression techniques will be compared in the case where  $X$  is a random variable.

### 1. Introduction

Given bivariate data  $\{(X_i, Y_i), i = 1, \dots, n\}$  one would like a technique for estimating the regression function  $r(x) = E(Y|X)$ , which is applicable for a wide range of functions,  $r(x)$ . Further, one would like a technique with good precision over the full range of values for  $X$ .

Techniques of quantile regression based on the approach developed by Parzen (1977), and discussed by Carmichael (1978), will be compared to each other and to Benedetti (1975) in the case where  $X$  is restricted to the unit interval. Further, the quantile regression techniques will be compared in the case where  $X$  is a random variable. Variations of each of the quantile regression estimators will also be considered in order to determine the most functionally appealing form of the estimators.

### 2. The Estimators

#### 2.1. $\hat{RQ}_1$

Given observations  $\{(X_i, Y_i), i = 1, \dots, n\}$  on random variables  $(X, Y)$  according to the model

$$Y_i = r(X_i) + \epsilon_i$$

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where  $r$  is an unknown function and the errors are i.i.d. random variables with zero mean and finite variance, the following versions of the estimator  $\hat{RQ}_1(u)$  are considered

$$\begin{aligned}
 (a) \quad & \frac{1}{nh(n)} \sum_{j=1}^n Y_{[j:n]} K\left(\frac{i-1}{n} - u\right) \\
 (b) \quad & \frac{1}{nh(n)} \left( \frac{Y_{[1:n]}}{2} K\left(\frac{0}{n} - u\right) + \frac{Y_{[n:n]}}{2} K\left(\frac{n-1}{n} - u\right) + \sum_{j=2}^{n-1} Y_{[j:n]} K\left(\frac{j-1}{n} - u\right) \right) \\
 (c) \quad & \left( \frac{Y_{[1:n]}}{2} K\left(\frac{0-u}{h(n)}\right) + \frac{Y_{[n:n]}}{2} K\left(\frac{n-1-u}{h(n)}\right) + \sum_{j=2}^{n-1} Y_{[j:n]} K\left(\frac{j-1-u}{h(n)}\right) \right) \\
 (d) \quad & \frac{\sum_{j=1}^n Y_{[j:n]} K\left(\frac{i-1}{n} - u\right)}{\sum_{j=1}^n K\left(\frac{i-1}{n} - u\right)}
 \end{aligned}$$

If one interpolates linearly between the successive points  $\{(j/n, Y_{[j:n]})\}$  one obtains a new estimator  $\hat{RQ}_2$ . The versions of this estimator are as follows:

$$\begin{aligned}
 (a) \quad & \frac{1}{nh(n)} \sum_{j=1}^n Y_{[j:n]} K\left(\frac{i-1}{n} - u\right) + \frac{1}{2nh(n)} \sum_{j=2}^n (Y_{[j-1:n]} - Y_{[j:n]}) K\left(\frac{n}{h(n)}\right) \\
 (b) \quad & \frac{\sum_{j=1}^n Y_{[j:n]} K\left(\frac{i-1}{n} - u\right)}{\sum_{j=1}^n K\left(\frac{i-1}{n} - u\right)} + \frac{1}{2nh(n)} \sum_{j=2}^n (Y_{[j-1:n]} - Y_{[j:n]}) K\left(\frac{n}{h(n)}\right)
 \end{aligned}$$

If instead of interpolating, one integrates a smoothed version of the first differences one obtains  $\hat{RQ}_3$ . The versions of this estimator looked at are:

$$\begin{aligned}
 & \int_{1/2}^u \hat{RQ}'_1(s) ds + \hat{RQ}_1(1/2) \\
 (a) \quad & \text{where } \hat{RQ}'_1(s) = \sum_{j=1}^{n-1} (Y_{[j+1:n]} - Y_{[j:n]}) \frac{1}{h(n)} K\left(\frac{i-1}{h(n)} - s\right) \text{ and} \\
 & \hat{RQ}_1(1/2) \text{ is as defined in 2.1(a)} \\
 K(Z) = & \begin{cases} \frac{15}{16} (1 - z^2)^2 & |z| \leq 1 \\ 0 & |z| > 1 \end{cases}
 \end{aligned}$$

$$(b) \text{ where } \hat{RQ}'_1(s) = \frac{n \sum_{j=1}^{n-1} (Y_{[j+1:n]} - Y_{[j:n]}) K\left(\frac{j-1}{n} - s\right)}{\sum_{j=1}^{n-1} K\left(\frac{j-1}{n} - s\right)}$$

and  $\hat{RQ}_1(1/2)$  is as defined in 2.1(d).

### 3. Comparison of the Estimators

One needs a way of comparing the estimators as to their efficiency over a range of functions. One can compare them for each function based on the criterion of minimizing mean integrated squared error (MISE). For fixed  $h(n)$ ,  $n$ , and  $r(x)$ , the MISE is given by

$$3.1. \text{ MISE} = E \int_0^1 (\hat{R}(x) - r(x))^2 dx = \int_0^1 \text{VAR}(\hat{R}(x)) dx + \int_0^1 (\text{Bias})^2 dx.$$

One will find in Appendix A the formulas used in the calculation of MISE for  $\hat{RQ}_1$  and  $\hat{RQ}_3$ , considered both in the case where  $X$  is fixed and where  $X$  is a random variable.

The MISE is also calculated over the interval (.25, .75) since the estimators perform most poorly near the endpoints.

#### $X$ Fixed

It was found that for a sample of size  $n$ , taking points of the form  $\frac{j-1}{n-1}$   $j = 1, \dots, n$ , rather than points of the form  $\frac{j}{n}$   $j = 1, \dots, n$ , reduced the MISE.

To accommodate this change in design the estimators must be slightly modified. Therefore the estimators  $\hat{RQ}_1(a)$  and  $\hat{RQ}_3(a)$  become

$$3.2. \quad \hat{RQ}_1 = \frac{1}{nh(n)} \sum_{j=1}^n Y_{[j:n]} K\left(\frac{j-1}{h(n)} - u\right) \text{ and}$$

$$3.3. \quad \hat{RQ}_3 = \int_{1/2}^u \hat{RQ}'_1(s) ds + \hat{RQ}_1(1/2) \quad \text{where}$$

$$\hat{RQ}'_1(s) = \sum_{j=1}^{n-1} (Y_{[j+1:n]} - Y_{[j:n]}) \frac{1}{h(n)} K\left(\frac{j-1}{h(n)} - s\right)$$

and  $\hat{RQ}_1(1/2)$  is as in 3.2.

Similar changes were also made in the other versions of each of the estimators.

Initially one would like to determine which of the four versions of  $\hat{RQ}_1$  is functionally the best. Tables 1 through 19 are the values calculated for the MISE for the different estimators, for the different models, around the optimum choice for the parameter  $M = 1/h(n)$ . A comparison of tables 1, 11, 12 and 13 indicates that the normalized versions of  $\hat{RQ}_1$  perform better than the un-normalized versions of  $\hat{RQ}_1$ . Therefore, one needs only look at the normalized versions,  $\hat{RQ}_1(c)$  and  $\hat{RQ}_1(d)$ .

Looking at Table 13 one finds that  $\hat{RQ1}(d)$  was better than  $\hat{RQ1}(c)$ , although little difference is found, and thus  $\hat{RQ1}(d)$  was chosen as the version to be looked at in more detail.

In preliminary simulations  $\hat{RQ2}$  did not perform substantially different from  $\hat{RQ1}$ .  $\hat{RQ2}(a)$  did not substantially change  $\hat{RQ1}(a)$  and therefore can be discarded just on the basis of the poor performance of  $\hat{RQ1}(a)$ . In these simulations  $\hat{RQ2}(b)$  looked very much like  $\hat{RQ1}(d)$  and in fact, based on a comparison of tables 1, 2, and 20 had higher integrated squared BIAS. From looking at the formula for  $\hat{RQ2}(b)$  there is good reason to suspect an increase in integrated variance. Therefore,  $\hat{RQ2}(b)$  will not be considered further.

It is also necessary to compare the two versions of  $\hat{RQ3}$  to determine which of these to study further.

One finds, by comparing Tables 1-3 with Tables 14-19, that normalizing  $\hat{RQ3}$  leads to a reduction in integrated squared bias but an increase in integrated variance. In fact, in the case where  $r(x)$  is linear, the bias of  $\hat{RQ3}(d)$  is exactly 0. This is a very desirable result. Still a choice is necessary to determine which merits further study.

One finds that the integrated squared bias and integrated variance are, for most cases, smaller for  $\hat{RQ1}(d)$  than  $\hat{RQ3}(a)$ . Therefore,

there is nothing to be gained in looking further at  $\hat{RQ3}(a)$ . Therefore, one needs only compare the two estimators  $\hat{RQ3}(b)$  and  $\hat{RQ1}(d)$  with each other and with the Benedetti estimator.

In preliminary simulations  $\hat{RQ2}$  did not perform substantially different from  $\hat{RQ1}$ .  $\hat{RQ2}(a)$  did not substantially change  $\hat{RQ1}(a)$  and therefore can be discarded just on the basis of the poor performance of  $\hat{RQ1}(a)$ . Therefore, one can only compare the MISE of her estimator with the ones discussed here.

Tables 1a, 2a, and 4a indicate that in all of the models tested  $\hat{RQ1}(d)$  performed at least as well as the Benedetti estimator and in fact, performed substantially better for the model  $Y = \sin(2\pi x)$ .

Now one needs to compare  $\hat{RQ1}(d)$  and  $\hat{RQ3}(b)$ . Tables 1 through 5 provide this comparison. One notes that  $\hat{RQ1}(d)$  has lower MISE in all the models. This is due to the increase in integrated variance in  $\hat{RQ3}(b)$ . In fact, the integrated squared bias of  $\hat{RQ3}(b)$  was better in the area of the optimum choice of  $M$  except for the model  $Y = \sin(2\pi x)$  for a sample size of 50. Clearly in all cases for a given value of  $M$ , the bias of  $\hat{RQ3}(b)$  is smaller than that of  $\hat{RQ1}(d)$ .

It should be noted that there was anticipation of an improvement by smoothing first differences. It is clear that in the normalized versions of the estimators the bias is reduced by smoothing first differences, but since the variance increases it is unclear whether  $\hat{RQ3}$  is an improvement.

If one compares Tables 11 and 17, one finds that when the estimators are not normalized, smoothing first differences represents a substantial improvement. In fact, by also looking at Table 14, one finds that  $\hat{RQ3}(a)$  has properties for a sample size of 20 equivalent to those of  $\hat{RQ1}(a)$  for a sample size of 50. These indeed represent the improvement that was anticipated.

#### X: Random Variable

For the case where  $X$  is a random variable only the estimators  $\hat{RQ1}(d)$  and  $\hat{RQ3}(b)$  are discussed since they were found to be the most appealing estimators for the case when  $X$  is fixed.

It is the belief here that a slight modification of the estimators will improve them. Let  $U_{(j)} = F(X_{(j)})$ ,  $j = 1, \dots, n$ . These are the order statistics for a uniform distribution. One can calculate their expectations and, in fact,

$$E[U_{(j)}] = \frac{j}{n+1}$$

Therefore, it is felt that evaluating the kernel at points of the form  $\frac{j}{n+1}$ ,  $j = 1, \dots, n$  will improve the estimators. Initial observations showed that this might in fact be true. Therefore, the estimators are now written as:

$$\begin{aligned} \hat{RQ1}(d) &= \frac{\sum_{j=1}^n Y_{[j:n]} K\left(\frac{j+1-u}{h(n)}\right)}{\sum_{j=1}^n K\left(\frac{n+1-u}{h(n)}\right)} \\ \text{and} \\ \hat{RQ3}(b) &= \hat{RQ1}(1/2) + (n+1) \int_{1/2}^u \frac{\sum_{j=1}^{n-1} (Y_{[j+1:n]} - Y_{[j:n]}) K\left(\frac{j+1-u}{h(n)}\right)}{\sum_{j=1}^{n-1} K\left(\frac{n+1-u}{h(n)}\right)} du \end{aligned}$$

These are the forms of the estimators which are discussed further.

Three models were looked at for two different sample sizes and a few different error structures.

From Tables 21 and 22 for the linear model  $Y = 1 - X$ , one finds as before that the bias of  $\hat{RQ3}(b)$  is substantially smaller than that of  $\hat{RQ1}(d)$ . Also as before, the variance for  $\hat{RQ1}(d)$  is smaller. However, one finds that provided the variance of the error is small enough the MISE is smaller for  $\hat{RQ3}(d)$  and even when the variance of the error increases the MISE of the two estimators are not very different.

Looking at Tables 23 and 24 one compares the performance of the estimators for the model  $Y = \sin(2\pi X)$ . One finds similar results.  $\hat{RQ3}(b)$  has substantially smaller bias and larger variance and again the MISE for the two estimators are roughly of the same order.

In Tables 25 and 26 the two estimators are compared for the model  $Y = \sqrt{\frac{1}{16} - \frac{X^2}{3}} (X + 4e^{-X^2/2})$ . For this model the results are slightly different. Again  $\hat{RQ3}(b)$  shows a very large reduction in bias and an increase in variance but for both sample sizes and both error structures the value of the MISE, at the optimum choice of  $M$ , for  $\hat{RQ3}(b)$  is substantially smaller than that of  $\hat{RQ1}(d)$ .

The results here clearly show  $\hat{RQ3}(b)$  has better bias properties than  $\hat{RQ1}(d)$  and also may indicate that the increase in variance may not be very significant. One might conclude that in fact  $\hat{RQ3}(b)$  represents an improvement of  $\hat{RQ1}(d)$ .

One might like to see this improvement so next these two estimators will be compared in simulations.

#### Simulations

Simulations were performed to see the differences in the estimators and to determine which performed better. Both the case of  $X$  fixed and  $X$  a random variable were studied.

The first model examined is  $Y = 3 + 2X$  when  $X$  is fixed. Figures 1, 1a, and 1b indicate the performance of the three estimators. It is clear that  $\hat{RQ2}(b)$  does not perform substantially different from  $\hat{RQ1}(d)$  and as indicated before it is in fact worse. This lends credence to the claim that  $\hat{RQ2}(b)$  is not an improvement and need not be considered further.

Looking only at the other two estimators, one sees that both perform well between .25 and .75 but at the endpoints  $\hat{RQ1}(d)$  veers away from the line while  $\hat{RQ3}(b)$  is barely discernable from the true line throughout.

When a random error is introduced in this model, one must re-evaluate the estimators. Here  $\hat{RQ1}(d)$  may be said to better estimate the true line but  $\hat{RQ3}(b)$  seems to follow the data more closely. This can be seen in the area near one. Here the data seems to swing upward and  $\hat{RQ3}(b)$  follows this pattern while  $\hat{RQ1}(d)$  stays below. (See figures 2 and 2a.)

Next the trigonometric function  $Y = \sin(2\pi X)$  with a random error introduced is looked at. The two estimators are not very different except near the endpoints. When  $X$  is near zero, it is unclear which performs better. When  $X$  is near one, however, the data follows the true regression more closely and  $\hat{RQ3}(b)$  better fits the data. (See figures 3 and 3a.)

The same function is examined in the case where  $X$  is distributed as a  $U(0, 1)$ . Again the differences in the estimators occur at the endpoints. When  $X$  is near zero one sees that the data is close to the true regression and  $\hat{RQ3}(b)$  estimates the regression better in this area. When  $X$  is near one, there is a clear upswing in the data which also is followed more closely by  $\hat{RQ3}(b)$ . (See figures 4 and 4a.)

Finally, the function  $Y = 1 - X$  with a random error is discussed when  $X$  is distributed  $N(0, 1)$ . Again, the estimators function alike except when  $X$  is near one. Here the data has an obvious downward trend which is clearly followed more closely by  $\hat{R}Q3(b)$  than  $RQ1(d)$ . (See figures 5 and 5a.)

#### 4. Conclusions

It was anticipated that the estimator  $\hat{R}Q3(b)$  would perform better especially near the endpoints. This would seem to be the case. Clearly, the estimators are similar in the middle but  $\hat{R}Q3(b)$ , while having more variability, substantially improves the estimation at the endpoints.

Also it was found that normalizing the estimators improved them. This was, however, more important for  $\hat{R}Q1$  than  $\hat{R}Q3$  but both were improved.

TABLE I

$$\begin{aligned} Y &= 3 + 2X \\ \sigma^2 &= .3 \\ N &= 20 \end{aligned}$$

$$M = \frac{1}{h(n)}$$

	1.5	2	2.5	3	3.5
<u>Bias Squared (0, 1)</u>					
$\hat{R}Q1(d)$	.0379	.0149	.0071	.0038	.0023
$\hat{R}Q3(b)$	0	0	0	0	0
<u>Variance* (0, 1)</u>					
$\hat{R}Q1(d)$	.0225	.0278	.0332	.0386	.0439
$\hat{R}Q3(b)$	.0453	.0441	.0464	.0504	.0552
<u>MISE* (0, 1)</u>					
$\hat{R}Q1(d)$	.0603	.0428	.0403	.0424	.0462
$\hat{R}Q3(b)$	.0453	.0441	.0464	.0504	.0552
<u>Bias Squared (-.25, .75)</u>					
$\hat{R}Q1(d)$	.0027	.0002	.0000	.0000	.0000
$\hat{R}Q3(b)$	0	0	0	0	0
<u>Variance* (-.25, .75)</u>					
$\hat{R}Q1(d)$	.0094	.0116	.0141	.0168	.0196
$\hat{R}Q3(b)$	.0103	.0123	.0156	.0193	.0232
<u>MISE* (-.25, .75)</u>					
$\hat{R}Q1(d)$	.0121	.0118	.0141	.0168	.0196
$\hat{R}Q3(b)$	.0103	.0123	.0156	.0193	.0232

\*These values are for  $\sigma^2 = .3$ . The integrated variance increases proportionally with  $\sigma^2$  and for a different value of  $\sigma^2$  must be adjusted. The MISE must also be adjusted accordingly.

$$\text{MISE} = \int_0^1 (\text{BLAS})^2 + \frac{\sigma^2}{3} \int_0^1 (\text{VAR})^2$$

TABLE 1a

Y = 3 + 2X

 $\sigma^2 = .3$ 

N = 50

$$M = \frac{1}{h(n)}$$

		2	2.5	3	3.5	4		
<u>Bias Squared (0, 1)</u>		.0176 0	.0088 0	.0050 0	.0030 0	.0020 0	<u>Bias Squared (0, 1)</u>	
<u>RQ1(d)</u>							<u>RQ1(d)</u>	
<u>RQ3(b)</u>							<u>RQ3(b)</u>	
<u>Variance* (0, 1)</u>		.0111 .0103	.0132 .0280	.0153 .0274	.0175 .0276	.0196 .0283	<u>Variance* (0, 1)</u>	
<u>RQ1(d)</u>							<u>RQ1(d)</u>	
<u>RQ3(b)</u>							<u>RQ3(b)</u>	
<u>MISE* (0, 1)</u>		.0287 .0303	.0220 .0280	.0203 .0274	.0205 .0276	.0216 .0283	<u>MISE* (0, 1)</u>	
<u>RQ1(d)</u>							<u>RQ1(d)</u>	
<u>RQ3(b)</u>							<u>RQ3(b)</u>	
<u>Bias Squared (.25, .75)</u>		.0003 0	.0000 0	.0000 0	.0000 0	.0000 0	<u>Bias Squared (.25, .75)</u>	
<u>RQ1(d)</u>							<u>RQ1(d)</u>	
<u>RQ3(b)</u>							<u>RQ3(b)</u>	
<u>Variance* (.25, .75)</u>		.0045 .0047	.0055 .0057	.0065 .0069	.0076 .0081	.0087 .0094	<u>Variance* (.25, .75)</u>	
<u>RQ1(d)</u>							<u>RQ1(d)</u>	
<u>RQ3(b)</u>							<u>RQ3(b)</u>	
<u>MISE* (.25, .75)</u>		.0049 .0047	.0055 .0057	.0065 .0069	.0076 .0081	.0087 .0094	<u>MISE* (.25, .75)</u>	
<u>RQ1(d)</u>							<u>RQ1(d)</u>	
<u>RQ3(b)</u>							<u>RQ3(b)</u>	

TABLE 2

Y = SIN(2πX)

 $\sigma^2 = .3$ 

N = 20

		2	2.5	3	3.5	4	5
<u>M = <math>\frac{1}{h(n)}</math></u>							
<u>Bias Squared (0, 1)</u>		.0448 .0448	.0243 .0243	.0292 .0171	.0196 .0155	.0095 .0095	
<u>RQ1(d)</u>							
<u>RQ3(b)</u>							
<u>Variance* (0, 1)</u>		.0464 .0464	.0504 .0504	.0439 .0552	.0493 .0608	.0601 .0601	
<u>RQ1(d)</u>							
<u>RQ3(b)</u>							
<u>MISE* (0, 1)</u>		.0912 .0912	.0833 .0747	.0732 .0723	.0690 .0763	.0697 .0763	
<u>RQ1(d)</u>							
<u>RQ3(b)</u>							
<u>Bias Squared (.25, .75)</u>		.0263 .0263	.0191 .0117	.0110 .0052	.0067 .0024	.0029 .0024	
<u>RQ1(d)</u>							
<u>RQ3(b)</u>							
<u>Variance* (.25, .75)</u>		.0156 .0156	.0168 .0193	.0196 .0232	.0224 .0229	.0280 .0280	
<u>RQ1(d)</u>							
<u>RQ3(b)</u>							
<u>MISE* (.25, .75)</u>		.0419 .0419	.0359 .0310	.0306 .0284	.0291 .0253	.0309 .0253	
<u>RQ1(d)</u>							
<u>RQ3(b)</u>							

Benedetti's Best

MISE*	.022
(0, 1)	.006
(.25, .75)	

TABLE 2a

$Y = \sin(2\pi X)$   
 $\sigma^2 = .3$   
 $N = 50$

$$M = \frac{1}{h(n)}$$

	3	3.5	4	5	6	8	
<u>Bias Squared (0, 1)</u>							
$\hat{R}Q1(d)$	.0258	.0137	.0224	.0117	.0067	.0026	
$\hat{R}Q3(b)$			.0078	.0040			
<u>Variance* (0, 1)</u>							
$\hat{R}Q1(d)$	.0274	.0276	.0196	.0239	.0282	.0367	
$\hat{R}Q3(b)$			.0283	.0386			
<u>MISE* (0, 1)</u>							
$\hat{R}Q1(d)$	.0532	.0413	.0420	.0356	.0348	.0393	
$\hat{R}Q3(b)$			.0361	.0426			
<u>Bias Squared (.25, .75)</u>							
$\hat{R}Q1(d)$	.0155	.0081	.0067	.0029	.0014	.0004	
$\hat{R}Q3(b)$			.0044	.0014			
<u>Variance* (.25, .75)</u>							
$\hat{R}Q1(d)$	.0069	.0081	.0087	.0109	.0130	.0174	
$\hat{R}Q3(b)$			.0094	.0119			
<u>MISE* (.25, .75)</u>							
$\hat{R}Q1(d)$	.0224	.0162	.0154	.0138	.0145	.0178	
$\hat{R}Q3(b)$			.0138	.0133			

TABLE 3

$Y = -X|2$   
 $\sigma^2 = .3$   
 $N = 20$

$$M = 1/h(n)$$

	1	1.5	2	2.5	3
<u>Bias Squared (0, 1)</u>					
$\hat{R}Q1(d)$	.0076	.0024	.0009	.0004	.0002
$\hat{R}Q3(b)$	0	0	0	0	0
<u>Variance* (0, 1)</u>					
$\hat{R}Q1(d)$	.0174	.0225	.0278	.0332	.0386
$\hat{R}Q3(b)$	.0492	.0453	.0441	.0464	.0504
<u>MISE* (0, 1)</u>					
$\hat{R}Q1(d)$	.0250	.0249	.0288	.0336	.0385
$\hat{R}Q3(b)$	.0492	.0453	.0441	.0464	.0504
<u>Bias Squared (.25, .75)</u>					
$\hat{R}Q1(d)$	.0010	.0002	.0000	.0000	.0000
$\hat{R}Q3(b)$	0	0	0	0	0
<u>Variance* (.25, .75)</u>					
$\hat{R}Q1(d)$	.0079	.0094	.0116	.0141	.0168
$\hat{R}Q3(b)$	.0111	.0103	.0123	.0156	.0193
<u>MISE* (.25, .75)</u>					
$\hat{R}Q1(d)$	.0089	.0096	.0116	.0141	.0165
$\hat{R}Q3(b)$	.0111	.0103	.0123	.0156	.0193

Benedetti's Best

MISE\*  
 $(0, 1)$   
 $(.25, .75)$

TABLE 3a

 $y = -x/2$   
 $\sigma^2 = .3$ 

N = 50

$M = \frac{1}{h(n)}$

	1	1.5	2	2.5	3	3.5	4
<u>Bias Squared (0, 1)</u>							
RQ1(d)	.0081	.0027	.0011	.0005	.0003	0	0
RQ3(b)		0	0	0	0	0	0
<u>Variance* (0, 1)</u>							
RQ1(d)	.0069	.0089	.0111	.0132	.0153	0.0276	.0283
RQ3(b)		.0303	.0280	.0274	.0274	.0283	
<u>MISE* (0, 1)</u>							
RQ1(d)	.0150	.0116	.0122	.0137	.0156	0.0276	.0283
RQ3(b)		.0303	.0280	.0274	.0274	.0283	
<u>Bias Squared (.25, .75)</u>							
RQ1(d)	.0011	.0002	.0000	.0000	.0000	0	0
RQ3(b)		0	0	0	0	0	0
<u>Variance* (.25, .75)</u>							
RQ1(d)	.0031	.0037	.0045	.0055	.0065	0.0081	.0094
RQ3(b)		.0047	.0057	.0069	.0069	.0081	.0094
<u>MISE* (.25, .75)</u>							
RQ1(d)	.0042	.0039	.0046	.0056	.0065	0.0081	.0094
RQ3(b)		.0047	.0057	.0069	.0069	.0081	.0094

TABLE 4

Y = 1 - X

 $\sigma^2 = .3$ 

N = 20

	1.5	2	2.5	3
<u>Bias Squared (0, 1)</u>				
RQ1(d)	.0095	.0037	.0018	.0010
RQ3(b)	0	0	0	0
<u>Variance* (0, 1)</u>				
RQ1(d)	.0225	.0278	.0332	.0386
RQ3(b)	.0453	.0441	.0464	.0504
<u>MISE* (0, 1)</u>				
RQ1(d)	.0320	.0316	.0350	.0395
RQ3(b)	.0453	.0441	.0464	.0504
<u>Bias Squared (.25, .75)</u>				
RQ1(d)	.0007	.0001	.0000	.0000
RQ3(b)	0	0	0	0
<u>Variance* (.25, .75)</u>				
RQ1(d)	.0094	.0116	.0141	.0168
RQ3(b)	.0103	.0123	.0156	.0193
<u>MISE* (.25, .75)</u>				
RQ1(d)	.0101	.0116	.0141	.0168
RQ3(b)	.0103	.0123	.0156	.0193

TABLE 4a

$Y = 1 - X$   
 $\sigma^2 = .3$   
 $N = 50$

	M = 1/h(n)					
	1.5	2	2.5	3	3.5	4
<u>Bias Squared (0, 1)</u>						
$\hat{R}Q1(d)$	.0107	.0044	.0022	.0012	.0008	
$\hat{R}Q3(b)$	0	0	0	0	0	
<u>Variance* (0, 1)</u>						
$\hat{R}Q1(d)$	.0089	.0111	.0132	.0153	.0175	
$\hat{R}Q3(b)$	.0303	.0280	.0274	.0276	.0283	
<u>MISE* (0, 1)</u>						
$\hat{R}Q1(d)$	.0196	.0155	.0154	.0166	.0182	
$\hat{R}Q3(b)$	.0303	.0280	.0274	.0276	.0183	
<u>Bias Squared (.25, .75)</u>						
$\hat{R}Q1(d)$	.0008	.0001	.0000	.0000	.0000	
$\hat{R}Q3(b)$	0	0	0	0	0	
<u>Variance* (.25, .75)</u>						
$\hat{R}Q1(d)$	.0037	.0045	.0055	.0065	.0076	
$\hat{R}Q3(b)$	.0047	.0057	.0057	.0069	.0081	.0094*
<u>MISE* (.25, .75)</u>						
$\hat{R}Q1(d)$	.0045	.0046	.0055	.0065	.0076	
$\hat{R}Q3(b)$	.0047	.0057	.0057	.0069	.0081	.0094

TABLE 5

$$Y = \sqrt{\frac{2}{7}}(X + 4e^{-X^2/2})$$

$\sigma^2 = .3$   
 $N = 20$

	M = 1/h(n)				
	1	1.5	2	2.5	3
<u>Bias Squared (0, 1)</u>					
$\hat{R}Q1(d)$	.0092	.0039	.0018	.0010	
$\hat{R}Q3(b)$	.0011	.0004	.0002	.0001	
<u>Variance* (0, 1)</u>					
$\hat{R}Q1(d)$	.0174	.0225	.0278	.0332	
$\hat{R}Q3(b)$	.0453	.0441	.0464	.0504	
<u>MISE* (0, 1)</u>					
$\hat{R}Q1(d)$	.0266	.0264	.0297	.0342	
$\hat{R}Q3(b)$	.0464	.0445	.0466	.0505	
<u>Bias Squared (.25, .75)</u>					
$\hat{R}Q1(d)$	.0024	.0010	.0004	.0002	
$\hat{R}Q3(b)$	.0009	.0003	.0001	.0000	
<u>Variance* (.25, .75)</u>					
$\hat{R}Q1(d)$	.0079	.0094	.0116	.0111	
$\hat{R}Q3(b)$	.0103	.0123	.0156	.0193	
<u>MISE* (.25, .75)</u>					
$\hat{R}Q1(d)$	.0103	.0104	.0120	.0143	
$\hat{R}Q3(b)$	.0112	.0126	.0157	.0193	

Benedetti's Best

MISE\*  
 $(0, 1)$   
 $(.25, .75)$

.016  
.005

TABLE 5a

$$Y = \sqrt{\frac{3}{7}}(X + 4e^{-x^2/2})$$

$\sigma^2 = .3$   
 $N = 50$

		M = 1/h(n)		
		1.5	2	2.5
Bias Squared	(0, 1)			
RQ1(d)	.0043	.0021	.0011	.0007
RQ3(b)		.0005	.0002	.0001
Variance*	(0, 1)			
RQ1(d)	.0089	.0111	.0132	.0153
RQ3(b)		.0303	.0280	.0274
MISE*	(0, 1)			
RQ1(d)	.0132	.0131	.0143	.0160
RQ3(b)		.0308	.0282	.0275

M = 1/h(n)

3  
3.5  
4  
5

RQ1(d)  
Bias Squared

Variance\* (.25, .75)

RQ1(d)  
RQ3(b)

Variance\* (.25, .75)

RQ1(d)  
RQ3(b)

Variance\* (.25, .75)

RQ1(d)  
RQ3(b)

		M = 1/h(n)		
		5	6	8
Bias Squared	(0, 1)			
RQ1(d)	.0010	.0004	.0002	.0001
RQ3(b)		.0003	.0002	.0001
Variance*	(0, 1)			
RQ1(d)	.0037	.0045	.0055	.0065
RQ3(b)		.0047	.0057	.0069
MISE*	(0, 1)			
RQ1(d)	.0047	.0049	.0057	.0066
RQ3(b)		.0050	.0059	.0070

TABLE 7

Y = SIN 2πX  
 $\sigma^2 = .3$   
N = 100

M = 1/h(n)

		M = 1/h(n)		
		5	6	8
Bias Squared	(0, 1)			
RQ1(d)	.0126	.0073	.0030	.0015
RQ3(b)		.0029	.0014	.0004
Variance*	(0, 1)			
RQ1(d)	.0119	.0141	.0183	.0226
RQ3(b)		.0054	.0065	.0086
MISE*	(0, 1)			
RQ1(d)	.0245	.0214	.0213	.0241
RQ3(b)		.0083	.0079	.0091

TABLE 6

Y = 3 + 2X  
 $\sigma^2 = .3$   
N = 100

		M = 1/h(n)		
		3	3.5	4
Bias Squared	(0, 1)			
RQ1(d)	.0054	.0033	.0022	.0011
RQ3(b)		.0000	.0000	.0000
Variance*	(0, 1)			
RQ1(d)	.0076	.0087	.0098	.0119
RQ3(b)		.0032	.0038	.0043
MISE*	(0, 1)			
RQ1(d)	.0130	.0121	.0120	.0130
RQ3(b)		.0032	.0038	.0043

TABLE 7

Y = SIN 2πX  
 $\sigma^2 = .3$   
N = 100

M = 1/h(n)

		M = 1/h(n)		
		5	6	8
Bias Squared	(0, 1)			
RQ1(d)	.0126	.0073	.0030	.0015
RQ3(b)		.0029	.0014	.0004
Variance*	(0, 1)			
RQ1(d)	.0119	.0141	.0183	.0226
RQ3(b)		.0054	.0065	.0086
MISE*	(0, 1)			
RQ1(d)	.0245	.0214	.0213	.0241
RQ3(b)		.0083	.0079	.0091

TABLE 8

$Y = 1 - X$   
 $\sigma^2 = .3$   
 $N = 100$

		M = 1/h(n)				
		RQI(d)			RQI(a)	
Bias Squared					Bias Squared	
(0, 1)	.0046	.0023	.0013	.0008	.0044	.0021
(.25, .75)	.0001	.0000	.0000	.0000	.0010	.0004
Variance*					(0, 1)	.0007
(0, 1)	.0055	.0066	.0076	.0087	(.25, .75)	.0001
(.25, .75)	.0023	.0027	.0032	.0038		
MISE*					(0, 1)	.0066
(0, 1)	.0102	.0089	.0090	.0096	(.25, .75)	.0076
(.25, .75)	.0024	.0027	.0032	.0038		.0032

TABLE 10

$Y = \sqrt{3/7} (X + 4e^{-x^2/2})$ ,  
 $\sigma^2 = .3$   
 $N = 100$

		M = 1/h(n)				
		RQI(d)			RQI(a)	
Bias Squared					Bias Squared	
(0, 1)	.0046	.0023	.0013	.0008	(0, 1)	.0007
(.25, .75)	.0001	.0000	.0000	.0000	(.25, .75)	.0001
Variance*					(0, 1)	.0007
(0, 1)	.0055	.0066	.0076	.0087	(.25, .75)	.0076
(.25, .75)	.0023	.0027	.0032	.0038		.0032
MISE*					(0, 1)	.0066
(0, 1)	.0088	.0077	.0078	.0084	(.25, .75)	.0032
(.25, .75)	.0028	.0026	.0029	.0033		

TABLE 11

$Y = 3 + 2X$   
 $\sigma^2 = .3$   
 $N = 50$

		M = 1/h(n)				
		RQI(d)			RQI(a)	
Bias Squared					Bias Squared	
(0, 1)	.0818	.0498	.0319	.0226	.0191	
(.25, .75)	.0032	.0032	.0032	.0033	.0051	
Variance*					(0, 1)	.0007
(0, 1)	.0497	.0665	.0834	.1002	.1171	
(.25, .75)	.0250	.0334	.0417	.0501	.0584	
MISE*					(0, 1)	.0007
(0, 1)	.1314	.1163	.1153	.1228	.13t2	
(.25, .75)	.0282	.0366	.0449	.0534	.0t35	

TABLE 12

		M = 1/h(n)		M = 1/h(n)	
		16	20	24	
		<u>RQI(b)</u>		<u>RQ3(a)</u>	
<u>Bias Squared</u>		<u>Bias Squared</u>		<u>Bias Squared</u>	
<u>Variance*</u>	(0, 1) (.25, .75)	.1072 .0000	.0875 .0000	.0768 .0002	.0871 .0380
<u>MISE*</u>	(0, 1) (.25, .75)	.0682 .0348	.0855 .0434	.1028 .0521	.0338 .0088
	(0, 1) (.25, .75)	.1754 .0348	.1730 .0434	.1796 .0523	.1210 .0468

TABLE 14

		M = 1/h(n)		M = 1/h(n)	
		16	20	24	3
		<u>RQI(b)</u>		<u>RQ3(a)</u>	
<u>Variance*</u>	(0, 1) (.25, .75)	.1072 .0000	.0875 .0000	.0768 .0000	.0465 .0202
<u>MISE*</u>	(0, 1) (.25, .75)	.0682 .0348	.0855 .0434	.1028 .0521	.0446 .0196

TABLE 13

		M = 1/h(n)		M = 1/h(n)	
		16	20	24	
		<u>RQI(b)</u>		<u>RQ3(a)</u>	
<u>Variance*</u>	(0, 1) (.25, .75)	.0715 .0711	.0888 .0883	.1061 .1055	.0445 .0144
<u>MISE*</u>	(0, 1) (.25, .75)	.0715 .0711	.0888 .0883	.1061 .1055	.0393 .0101

TABLE 15

		M = 1/h(n)		M = 1/h(n)	
		16	20	24	5
		<u>RQI(b)</u>		<u>RQ3(a)</u>	
<u>Variance*</u>	(0, 1) (.25, .75)	.0715 .0711	.0888 .0883	.1061 .1055	.0445 .0144
<u>MISE*</u>	(0, 1) (.25, .75)	.0715 .0711	.0888 .0883	.1061 .1055	.0393 .0101
	(0, 1) (.25, .75)	.0348 .0348	.0434 .0434	.0521 .0521	.0458 .0206
	(0, 1) (.25, .75)	.0348 .0348	.0434 .0434	.0521 .0521	.0502 .0261
	(0, 1) (.25, .75)	.0348 .0348	.0434 .0434	.0521 .0521	.0599 .0223
	(0, 1) (.25, .75)	.0348 .0348	.0434 .0434	.0521 .0521	.0599 .0223

TABLE 16

		M = 1/h(n)			M = 1/h(n)		
		$\hat{RQ3}(a)$			$\hat{RQ3}(a)$		
		<u>Bias Squared</u>			<u>Bias Squared</u>		
$\sigma^2$	.3	(0, 1)	.0009	.0004	.0003	.0003	.0065
N	20	(.25, .75)	.0002	.0001	.0001	.0001	.0009
		<u>Variance*</u>			(0, 1)	.0183	.0112
		(.25, .75)			(.25, .75)	.0072	.0034
		<u>MISE*</u>			(0, 1)	.0245	.0273
		(.25, .75)			(.25, .75)	.0084	.0126
		<u>MISE*</u>			(0, 1)	.0427	.0385
		(.25, .75)			(.25, .75)	.0156	.0139
							.0145
							.0179

TABLE 18

		M = 1/h(n)			M = 1/h(n)		
		$\hat{RQ3}(a)$			$\hat{RQ3}(a)$		
		<u>Bias Squared</u>			<u>Bias Squared</u>		
$\sigma^2$	.3	(0, 1)	.0338	.0362	.0384	.0418	.0381
N	50	(.25, .75)	.0088	.0105	.0128	.0153	.0169
		<u>Variance*</u>			(0, 1)	.0245	.0273
		(.25, .75)			(.25, .75)	.0084	.0126
		<u>MISE*</u>			(0, 1)	.0427	.0385
		(.25, .75)			(.25, .75)	.0156	.0139
							.0145
							.0179

TABLE 19

		M = 1/h(n)			M = 1/h(n)		
		$\hat{RQ3}(a)$			$\hat{RQ3}(a)$		
		<u>Bias Squared</u>			<u>Bias Squared</u>		
$\sigma^2$	.3	(0, 1)	.0309	.0303	.0307	.0315	.0235
N	50	(.25, .75)	.0085	.0094	.0105	.0115	.0073
		<u>Variance*</u>			(0, 1)	.0235	.0227
		(.25, .75)			(.25, .75)	.0045	.0053
		<u>MISE*</u>			(0, 1)	.0226	.0227
		(.25, .75)			(.25, .75)	.0063	.0073

TABLE 17

		M = 1/h(n)			M = 1/h(n)		
		$\hat{RQ3}(a)$			$\hat{RQ3}(a)$		
		<u>Bias Squared</u>			<u>Bias Squared</u>		
$\sigma^2$	.3	(0, 1)	.0085	.0077	.0073	.0070	.0001
N	50	(.25, .75)	.0032	.0032	.0032	.0032	.0000
		<u>Variance*</u>			(0, 1)	.0002	.0001
		(.25, .75)			(.25, .75)	.0000	.0000
		<u>MISE*</u>			(0, 1)	.0232	.0225
		(.25, .75)			(.25, .75)	.0045	.0053

TABLE 20

N = 50

M = 1/h(n)

<u>RQ2(b)</u>	2	3	4	5
<u>Y = 3 + 2X :</u>				
<u>Bias Squared</u>				
(0, 1)	.0180	.0053	.0024	.0014
(.25, .75)	.0006	.0002	.0002	.0002
<u>Y = SIN(2πX)</u>				
<u>Bias Squared</u>				
(0, 1)	.1157	.0490	.0240	.0135
(.25, .75)	.0629	.0199	.0077	.0039

	M = 1/h(n)				
	1.5	2	2.5	3	3.5
<u>Bias Squared (<math>\frac{1}{302}, \frac{301}{302}</math>)</u>					
RQ1(d)	.0379	.0241	.0813	.0510	.0453
RQ3(b)			.0204	.0183	
<u>ε ~ N(0, 3)</u>					
<u>Variance (<math>\frac{1}{302}, \frac{301}{302}</math>)</u>					
RQ1(d)	.1396	.1379	.1072	.1159	.1243
RQ3(b)			.1395	.1433	
<u>MISE (<math>\frac{1}{302}, \frac{301}{302}</math>)</u>					
RQ1(d)	.1775	.1620	.1885	.1769	.1726
RQ3(b)			.1599	.1616	
<u>ε ~ N(0, 1)</u>					
<u>Variance (<math>\frac{1}{302}, \frac{301}{302}</math>)</u>					
RQ1(d)	.2373	.2338	.1736	.1936	.2137
RQ3(b)			.2390	.2496	
<u>MISE (<math>\frac{1}{302}, \frac{301}{302}</math>)</u>					
RQ1(d)	.2752	.2579	.2545	.2536	.2620
RQ3(b)			.2679		
<u>ε ~ Double Exp. σ² = .3</u>					
<u>Variance (<math>\frac{1}{302}, \frac{301}{302}</math>)</u>					
RQ1(d)	.1416	.1442	.1167	.1267	.1358
RQ3(b)			.1487	.1550	
<u>MISE (<math>\frac{1}{302}, \frac{301}{302}</math>)</u>					
RQ1(d)	.1795	.1683	.1980	.1877	.1841
RQ3(b)			.1691	.1733	

TABLE 22

	M	2.5	3	3.5	4	5	6
<u>Bias Squared (<math>\frac{1}{302}, \frac{301}{302}</math>)</u>							
$\hat{R}Q1(d)$		.0198	.0144	.0437	.0353	.0283	.0221
$\hat{R}Q3(b)$				.0110	.0102		
$\epsilon \sim N(0, .3)$							
<u>Variance (<math>\frac{1}{302}, \frac{301}{302}</math>)</u>		.0754	.0755	.0604	.0638	.0702	.0763
$\hat{R}Q1(d)$				.0766	.0782		
$\hat{R}Q3(b)$							
<u>MISE (<math>\frac{1}{302}, \frac{301}{302}</math>)</u>		.0952	.0899	.1041	.0991	.0985	.0984
$\hat{R}Q1(d)$				.0876	.0884		
$\hat{R}Q3(b)$							
$\epsilon \sim N(0, 1)$							
<u>Variance (<math>\frac{1}{302}, \frac{301}{302}</math>)</u>		.1398	.1391	.1030	.1113	.1277	.1440
$\hat{R}Q1(d)$				.1414	.1454		
$\hat{R}Q3(b)$							
<u>MISE (<math>\frac{1}{302}, \frac{301}{302}</math>)</u>		.1596	.1535	.1467	.1466	.1560	.1661
$\hat{R}Q1(d)$				.1524	.1556		
$\hat{R}Q3(b)$							

TABLE 23

	M	2.5	3	3.5	4	5
<u>Bias Squared (0, 1)</u>						
$\hat{R}Q1(d)$		.0805	.0772	.0619	.0515	.0385
$\hat{R}Q3(b)$			.0557	.0445	.0373	
$\epsilon \sim N(0, .3)$						
<u>Variance (0, 1)</u>		.1066	.1002	.1140	.1265	.1481
$\hat{R}Q1(d)$			.1220	.1370	.1513	
$\hat{R}Q3(b)$						
<u>MISE (0, 1)</u>		.1871	.1774	.1759	.1780	.1866
$\hat{R}Q1(d)$			.1777	.1815	.1886	
$\hat{R}Q3(b)$						
$\epsilon \sim U(-1, 1)$						
<u>Variance (0, 1)</u>		.1244	.1025	.1157	.1278	.1493
$\hat{R}Q1(d)$			.1380	.1506	.1623	
$\hat{R}Q3(b)$						
<u>MISE (0, 1)</u>		.2049	.1797	.1776	.1793	.1878
$\hat{R}Q1(d)$			.1937	.1951	.1996	
$\hat{R}Q3(b)$						

TABLE 24

$Y = \sin(2\pi X)$   
 $X \sim U(0, 1)$   
 $N = 50$

 $M = 1/h(n)$ 

	3.5	4	5	6	8	
<u>Bias Squared (0, 1)</u>						
$\hat{R}Q1(d)$	.0256	.0354	.0228	.0159	.0094	
$\hat{R}Q3(b)$		.0180	.0127	.0103		
$\epsilon \sim N(0, .3)$						
<u>Variance (0, 1)</u>						
$\hat{R}Q1(d)$	.0545	.0497	.0585	.0658	.0782	
$\hat{R}Q3(b)$		.0583	.0621	.0771		
<u>MISE (0, 1)</u>						
$\hat{R}Q1(d)$	.0801	.0851	.0813	.0817	.0876	
$\hat{R}Q3(b)$		.0763	.0748	.0884		
$\epsilon \sim U(-1, 1)$						
<u>Variance (0, 1)</u>						
$\hat{R}Q1(d)$	.0555	.0452	.0539	.0615	.0747	
$\hat{R}Q3(b)$		.0582	.0612	.0815		
<u>MISE (0, 1)</u>						
$\hat{R}Q1(d)$	.0811	.0806	.0767	.0774	.0841	
$\hat{R}Q3(b)$		.0762	.0739	.0915		

TABLE 25

$Y = \sqrt{\frac{\sqrt{3}}{16-7\sqrt{3}}} (X + 4e^{-x^2/2})$   
 $X \sim N(0, 1)$   
 $N = 20$

 $M = 1/h(n)$ 

	2.5	3	3.5	4	5	6	
<u>Bias Squared (<math>\frac{1}{302}, \frac{301}{302}</math>)</u>							
$\hat{R}Q1(d)$	.0941	.1766	.1428	.1209	.0916	.0738	
$\hat{R}Q3(b)$		.0731	.0619	.0551	.0473		
$\epsilon \sim N(0, .3)$							
<u>Variance (<math>\frac{1}{302}, \frac{301}{302}</math>)</u>							
$\hat{R}Q1(d)$	.1951	.1844	.1985	.2105	.2313	.2495	
$\hat{R}Q3(b)$		.2064	.2172	.2280	.2485		
<u>MISE (<math>\frac{1}{302}, \frac{301}{302}</math>)</u>							
$\hat{R}Q1(d)$	.2892	.3610	.3423	.3314	.3229	.3233	
$\hat{R}Q3(b)$		.2795	.2791	.2831	.2958		
$\epsilon \sim N(0, 1)$							
<u>Variance (<math>\frac{1}{302}, \frac{301}{302}</math>)</u>							
$\hat{R}Q1(d)$	.3009	.2730	.2993	.3235	.3693	.4128	
$\hat{R}Q3(b)$		.3198	.3379	.3590	.4037		
<u>MISE (<math>\frac{1}{302}, \frac{301}{302}</math>)</u>							
$\hat{R}Q1(d)$	.3950	.4496	.4431	.4444	.4609	.4866	
$\hat{R}Q3(b)$		.3920	.3998	.4141	.4510		

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TABLE 26

$$Y = \sqrt{\frac{3}{16 - 7\sqrt{3}}} (X + 4e^{-x^2/4})$$

$X \sim N(0, 1)$

$N = 50$

$M = 1/h(n)$

<u>Bias Squared</u> ( $\frac{1}{302}, \frac{301}{302}$ )	
$\hat{R}Q1(d)$	.0297
$RQ3(b)$	.0203
<u>Variance</u> ( $\frac{1}{302}, \frac{301}{302}$ )	
$\hat{R}Q1(d)$	.0829
$RQ3(b)$	.0866
<u>MSE</u> ( $\frac{1}{302}, \frac{301}{302}$ )	
$\hat{R}Q1(d)$	.1126
$RQ3(b)$	.1069
$\epsilon \sim N(0, 1)$	
<u>Variance</u> ( $\frac{1}{302}, \frac{301}{302}$ )	
$\hat{R}Q1(d)$	.1424
$RQ3(b)$	.1466
<u>MSE</u> ( $\frac{1}{302}, \frac{301}{302}$ )	
$\hat{R}Q1(d)$	.1721
$RQ3(b)$	.1669

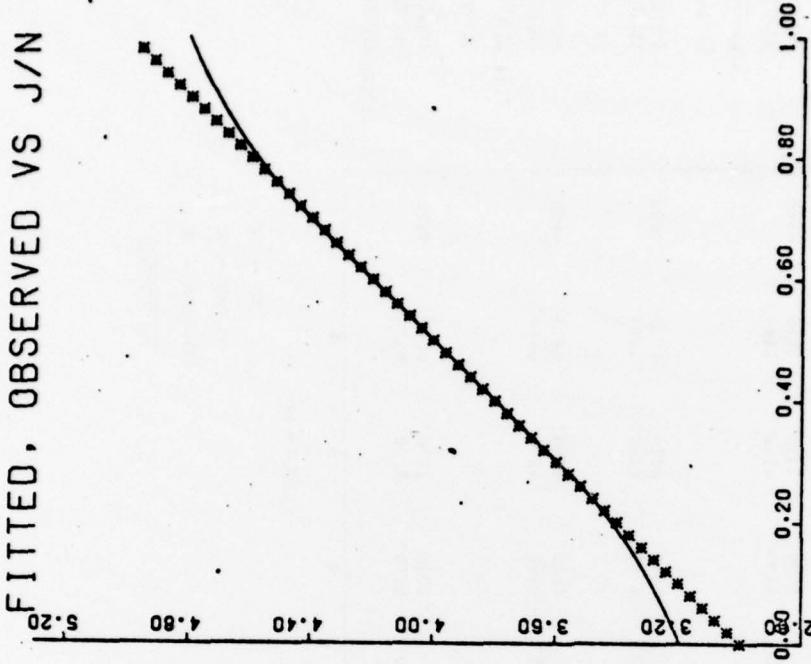


Figure 1.

$\hat{R}Q1(d)$

$Y = 3 + 2X$   
X : fixed  
N : 50  
M : 3

KEY:

\* DATA  
X TRUE REGRESSION  
ESTIMATOR

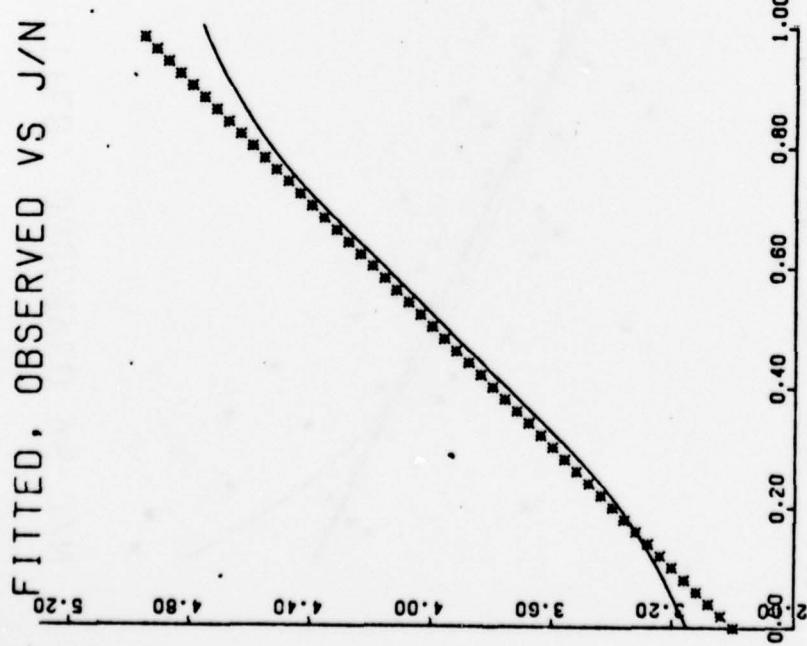


Figure 1a.  
RQ2(b)  
 $Y = 3 + 2X$   
X : fixed  
N = 50  
M = 3

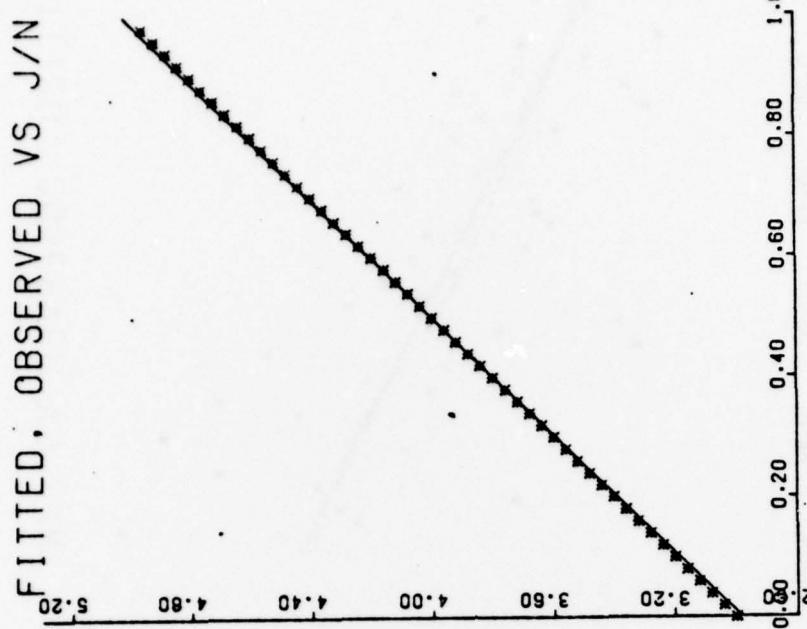


Figure 1b.  
RQ3(d)  
 $Y = 3 + 2X$   
X : fixed  
N = 50  
M = 3

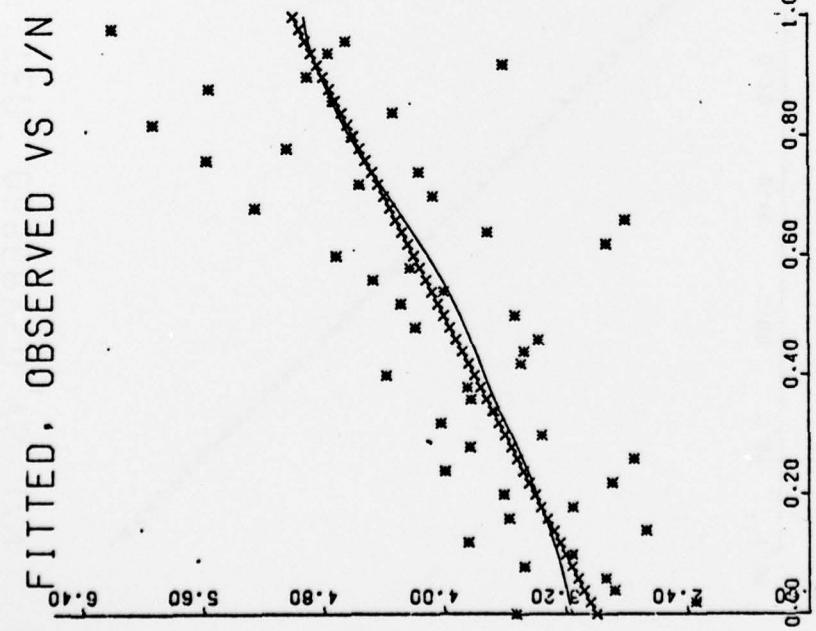


Figure 2.  
RQ1 (d)  
 $Y = 3 + 2X + \epsilon$   
 $\epsilon \sim N(0, .3)$   
X : fixed  
N = 50  
M = .3

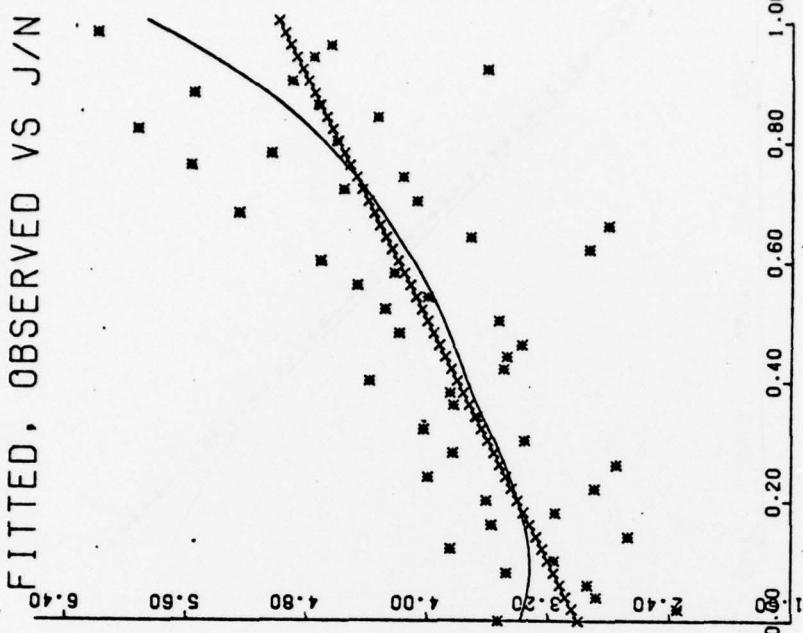


Figure 2a.  
RQ3 (b)  
 $Y = 3 + 2X + \epsilon$   
 $\epsilon \sim N(0, .3)$   
X : fixed  
N = 50  
M = .3

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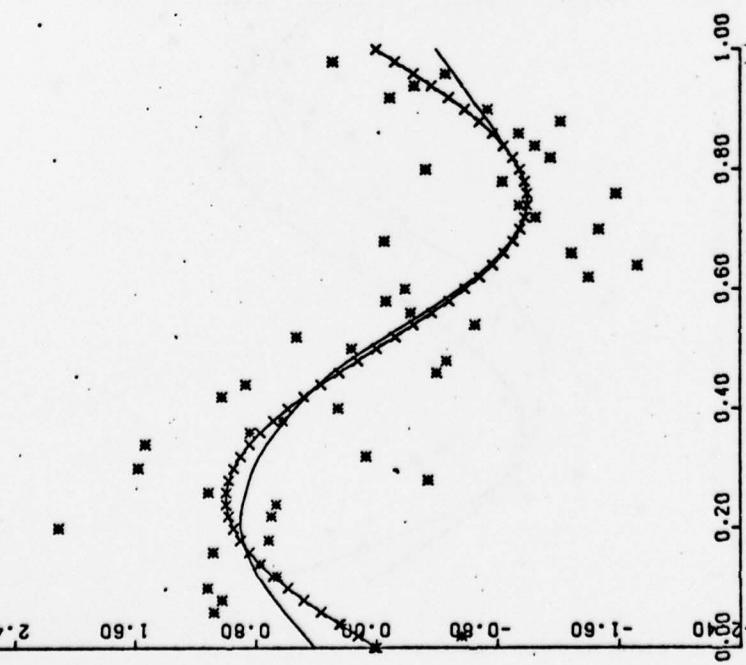


Figure 3.

RQ1 (d)  
 $Y = \sin(2\pi X) + \epsilon$   
 $\epsilon \sim N(0, .3)$   
X fixed  
N = 50  
M = 5

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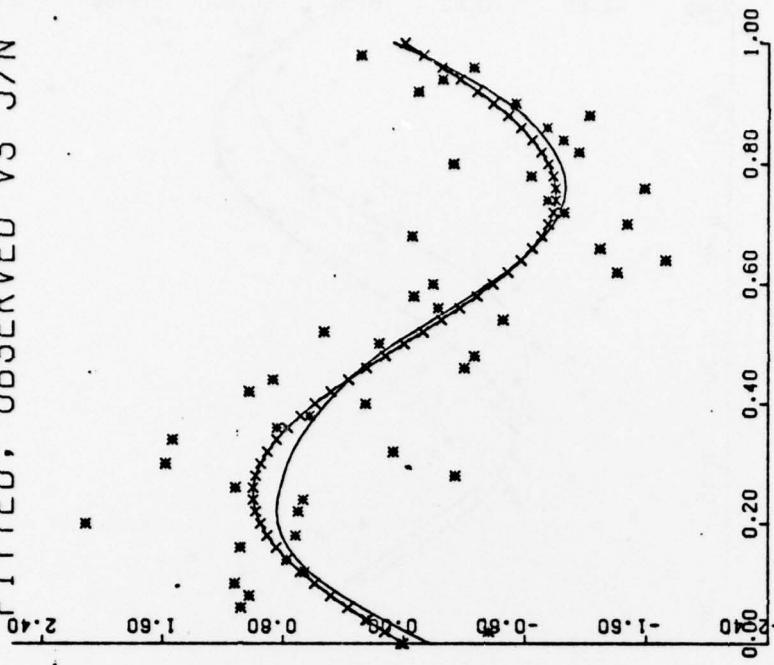


Figure 3a.

RQ3 (b)  
 $Y = \sin(2\pi X) + \epsilon$   
 $\epsilon \sim N(0, .3)$   
X fixed  
N = 50  
M = 5

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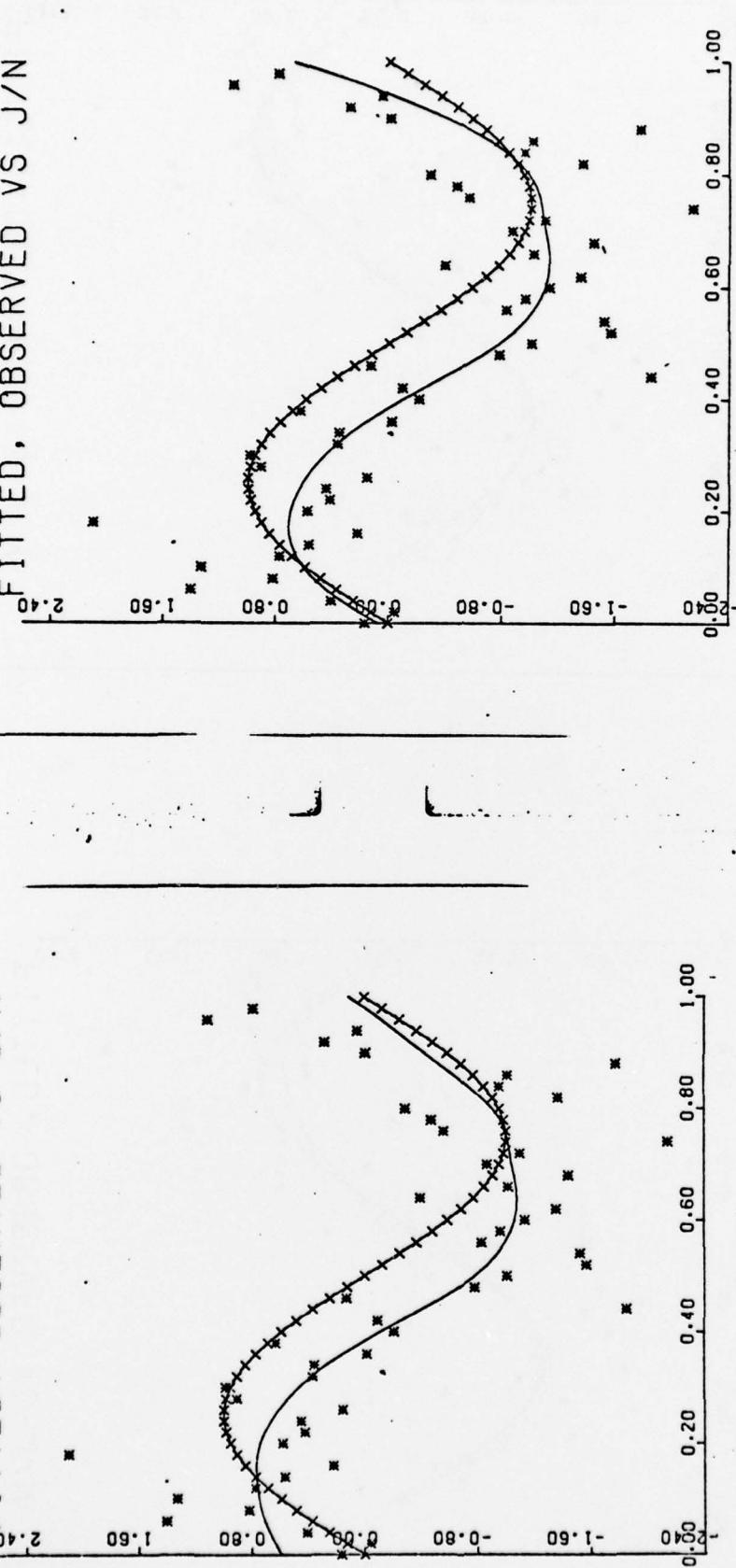


Figure 4.

RQ3 (b)  
 $Y = \sin(2\pi X) + \epsilon$   
 $\epsilon \sim N(0, .3)$   
 $X \sim U(0, 1)$   
 $N = 50$   
 $M = 5$

FITTED, OBSERVED VS J/N

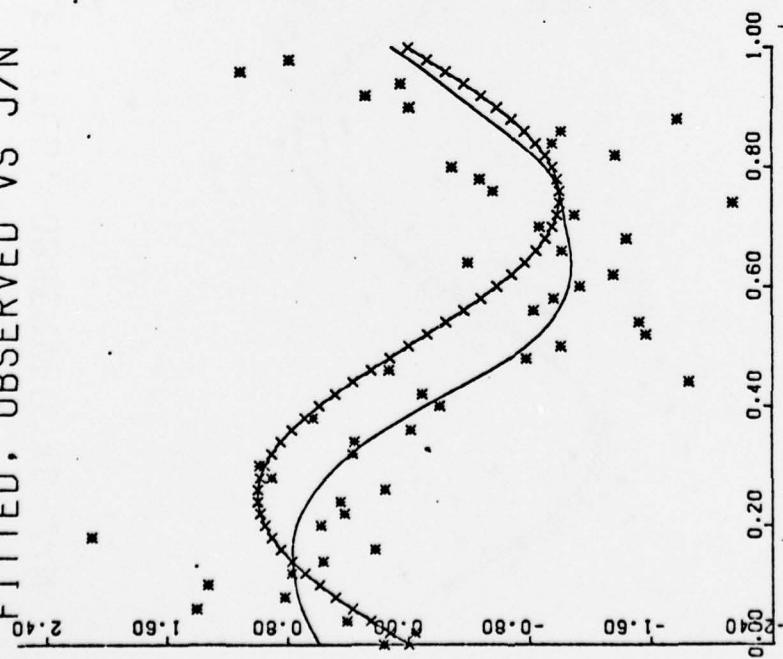


Figure 4.

RQ1 (d)  
 $Y = \sin(2\pi X) + \epsilon$   
 $\epsilon \sim N(0, .3)$   
 $X \sim U(0, 1)$   
 $N = 50$   
 $M = 5$

FITTED, OBSERVED VS J/N

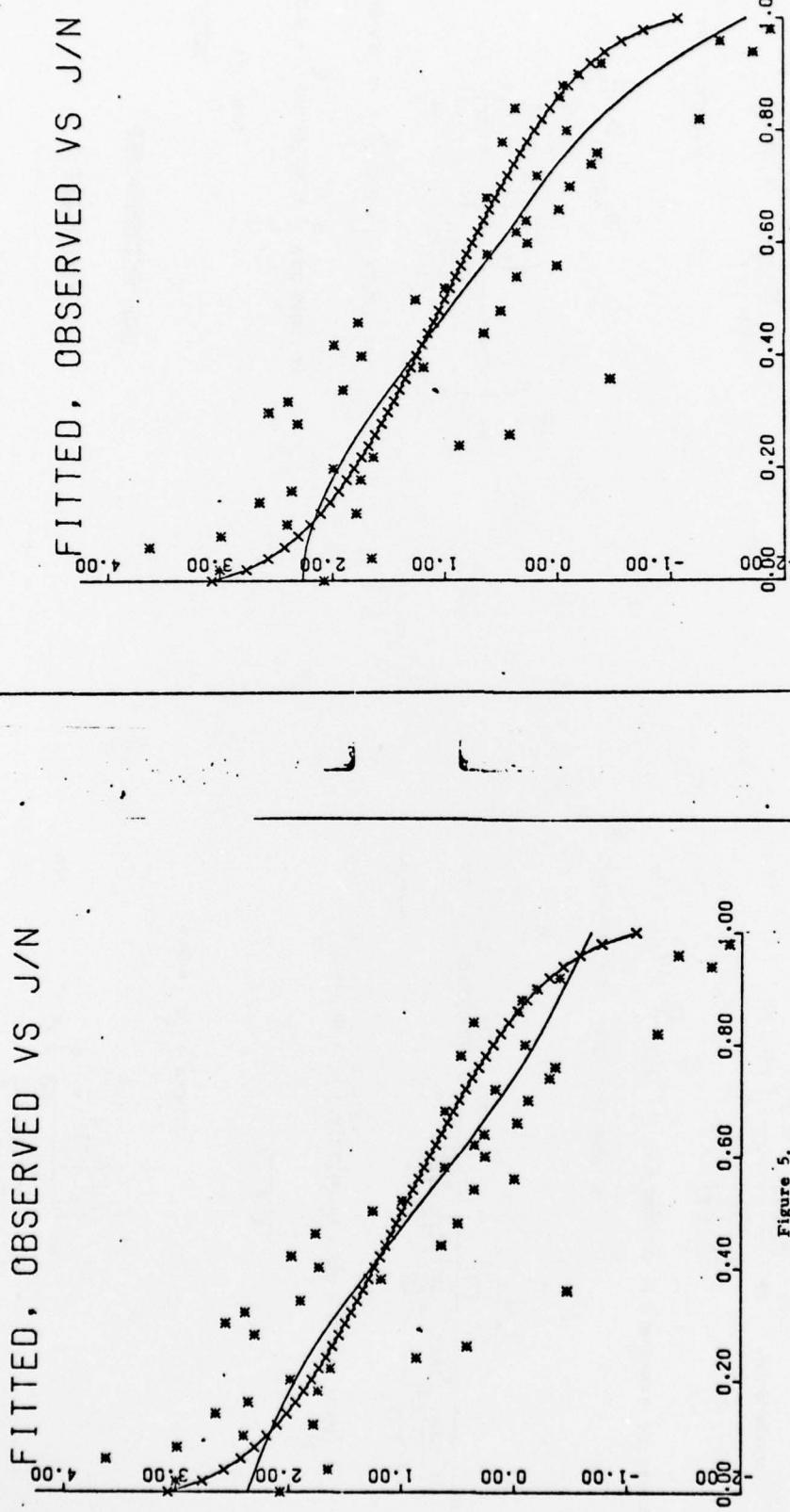


Figure 5.

RQ1 (d)  
 $Y = 1 - X + \epsilon$   
 $X \sim N(0, 1)$   
 $\epsilon \sim N(0, .3)$   
N = 50  
M = 3

FITTED, OBSERVED VS J/N

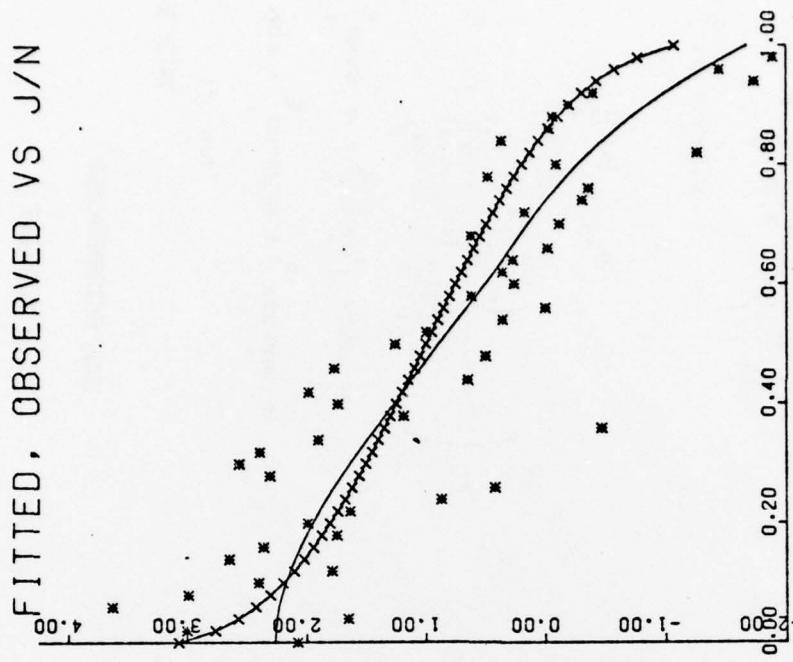


Figure 5a.

RQ3 (b)  
 $Y = 1 - X + \epsilon$   
 $X \sim N(0, 1)$   
N = 50  
M = 3  
 $\epsilon \sim N(0, .3)$

APPENDIX A  
FORMULAS FOR MSE

A. X Fixed

1.  $\hat{RQ}_1$

$$\text{MSE} = \int_0^1 (\text{BIAS})^2 dx + \int_0^1 \text{VAR}(R(x)) dx$$

$$\int_0^1 (\text{BIAS})^2 dx = \int_0^1 (E[\hat{RQ}_1] - rQ(u))^2 du$$

$$A1. \quad = \int_0^1 \left( \frac{n}{\sum_{j=1}^n E[Y_{[j:n]}]} K\left(\frac{i-1}{h(n)}\right) - rQ(u) \right)^2 du$$

$$= \left( \frac{n}{\sum_{j=1}^n K\left(\frac{i-1}{h(n)}\right)} \right)^2$$

where  $E[Y_{[j:n]}] = r(x_{(j)})$

where  $\sigma^2 = \text{VAR}(\epsilon)$

2.  $\hat{RQ}_3$

$$\int_0^1 (\text{BIAS})^2 dx = \int_0^1 \left( \int_{1/2}^u E[\hat{RQ}'_1(s)] ds + E[\hat{RQ}_1(1/2)] - rQ(u) \right)^2 du$$

where:

$$A3. \quad \int_{1/2}^u E[\hat{RQ}'_1(s)] ds = \frac{u}{\sum_{j=1}^{n-1} K\left(\frac{j-1}{h(n)}\right)} \int_{1/2}^u \frac{\sum_{j=1}^{n-1} E[Y_{[j+1:n]}] E[Y_{[j:n]}]}{\sum_{j=1}^{n-1} K\left(\frac{j-1}{h(n)}\right)} ds$$

but when  $r(x)$  is a linear function.

$$d_j = E[Y_{[j+1:n]}] - E[Y_{[j:n]}]$$

is a constant  $d$ ; therefore A3 becomes

$$A4. \quad \frac{(n-1)d}{\sum_{j=1}^{n-1} K\left(\frac{j-1}{h(n)}\right)} \int_{1/2}^u \frac{u}{\sum_{j=1}^{n-1} K\left(\frac{j-1}{h(n)}\right)} ds$$

which equals

$$\int_0^1 \frac{\sum_{j=1}^n \text{VAR}(Y_{[j:n]}) K^2\left(\frac{i-1}{h(n)}\right)}{\left(\sum_{j=1}^n K\left(\frac{i-1}{h(n)}\right)\right)^2} du$$

Since the  $\epsilon$ 's are uncorrelated

$$A5. \quad (n-1)d(u-1/2)$$

where

$$A8. \quad E[Y_{[j:n]}] = \int_0^1 rQ(u) n \binom{n-1}{j-1} u^{j-1} (1-u)^{n-j} du .$$

This was the formula used for calculating the squared bias. As before the trapezoidal rule was used in estimating the integral. Clearly for some distributions of  $X$  the integral can not be calculated over the whole range of  $u$ . Since 300 intervals were being used in calculating the integral a natural alternative is to calculate the integral over the region  $(\frac{1}{302}, \frac{301}{302})$ . This was used where necessary.

The formula for the variance was computationally difficult and thus the variance was calculated by 100 simulations on each of the models using the trapezoidal rule as before.

RQ3

$$\int_0^1 (BLAS)^2 = \int_0^1 \left( \int_{1/2}^u E[\hat{RQ}_1'(s)] ds + E[\hat{RQ}_1(1/2)] - rQ(u) \right)^2 du$$

where

$$\begin{aligned} & \sum_{j=1}^{n-1} \left( E[Y_{[j+1:n]}] - E[Y_{[j:n]}] \right) K \left( \frac{j}{h(n)} \cdot s \right) \\ & \int_{1/2}^u E[\hat{RQ}_1'(s)] ds = \int_{1/2}^u (n+1) \sum_{j=1}^{n-1} K \left( \frac{j}{h(n)} \right) \end{aligned}$$

$$\begin{aligned} & = (n+1) \sum_{j=1}^{n-1} \left( E[Y_{[j+1:n]}] - E[Y_{[j:n]}] \right) \int_0^1 K \left( \frac{j}{h(n)} \right) \\ & \quad \sum_{j=1}^{n-1} K \left( \frac{j}{h(n)} \right) ds \end{aligned}$$

$$\begin{aligned} & \text{where } E[Y_{[j:n]}] \text{ is as given in A8} . \\ & \text{Again the variance is computationally difficult so simulations were used here also.} \end{aligned}$$

$$E[\hat{R}Q_1(1/2)] = \frac{\sum_{j=1}^n E[Y_{[j:n]}] K\left(\frac{j-1}{n-1}, \frac{1}{2}\right)}{\sum_{j=1}^n K\left(\frac{j-1}{n-1}, \frac{1}{2}\right)}$$

$$\begin{aligned} \text{VAR}(\hat{R}Q_3) &= \int_0^1 \left[ \text{VAR}\left(\int_{1/2}^u \hat{R}Q'_1(s) ds\right) + \text{VAR}(\hat{R}Q_1(1/2)) + \right. \\ &\quad \left. 2 \text{COV}\left(\int_{1/2}^u \hat{R}Q'_1(s) ds, \hat{R}Q_1(1/2)\right) \right] du \end{aligned}$$

$$\text{COV}\left(\int_{1/2}^u \hat{R}Q'_1(s) ds, \hat{R}Q_1(1/2)\right) =$$

$$\begin{aligned} & (n-1) \sigma^2 \sum_{j=1}^{n-1} \int_{1/2}^u K\left(\frac{j-1}{n-1}, \frac{s}{h(n)}\right) ds \cdot K\left(\frac{n-j}{n-1}, \frac{1}{2}\right) \\ & \quad \times \sum_{j=1}^{n-1} K\left(\frac{j-1}{n-1}, \frac{s}{h(n)}\right) \sum_{j=1}^{n-1} K\left(\frac{n-j}{n-1}, \frac{1}{2}\right) \end{aligned}$$

Each of these integrals was evaluated using the trapezoidal rule.

The interval (0, 1) was divided into 300 equal segments for purposes of this evaluation.

The integrals in A6 and A7 were done using IMSL subroutine DCADRE with relative error set to be .01. Since the values of these integrals are small, the error will be very small and an estimate of the error from the subroutine itself was less than .00001.

$$\begin{aligned} \text{VAR}\left(\int_{1/2}^u \hat{R}Q'_1(s) ds\right) &= \int_{1/2}^u \int_{1/2}^u \text{COV}(\hat{R}Q'_1(s), \hat{R}Q'_1(t)) ds dt \\ &= (n-1) \sigma^2 \sum_{j=1}^{n-1} \int_{1/2}^u \frac{K\left(\frac{j-1}{n-1}, \frac{s}{h(n)}\right)}{\sum_{j=1}^{n-1} K\left(\frac{j-1}{n-1}, \frac{s}{h(n)}\right)} ds \\ &\quad \times \int_{1/2}^u \frac{K\left(\frac{j-1}{n-1}, \frac{t}{h(n)}\right)}{2K\left(\frac{j-1}{n-1}, \frac{t}{h(n)}\right)} \frac{K\left(\frac{j-2}{n-1}, \frac{t}{h(n)}\right)}{\sum_{j=1}^{n-1} K\left(\frac{j-1}{n-1}, \frac{t}{h(n)}\right)} dt \end{aligned}$$

where

$$\begin{aligned} \text{COV}(\hat{R}Q'_1(s), \hat{R}Q'_1(t)) &= \int_{1/2}^u \int_{1/2}^t \text{COV}(\hat{R}Q'_1(s), \hat{R}Q'_1(t)) ds dt \\ &= (n-1) \sigma^2 \sum_{j=1}^{n-1} \int_{1/2}^u \frac{K\left(\frac{j-1}{n-1}, \frac{s}{h(n)}\right)}{\sum_{j=1}^{n-1} K\left(\frac{j-1}{n-1}, \frac{s}{h(n)}\right)} ds \end{aligned}$$

and

$$\begin{aligned} \int_0^1 (\text{BLAS})^2 dx &= \int_0^1 (E[\hat{R}Q_1] - rQ(u))^2 du \\ &= \int_0^1 \left( \sum_{j=1}^n E[Y_{[j:n]}] K\left(\frac{j-1}{n-1}, \frac{u}{h(n)}\right) - rQ(u) \right)^2 du \end{aligned}$$

X: RANDOM VARIABLE

RQ1

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